Lecture 8: Probability Distributions

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Random Variables

- A sample space is a set of outcomes from an experiment. We denote this by S.
- A random variable is a function which maps outcomes into the real line. It is given by x : S→R.
- Each element in the sample space has an associated probability and such probabilities sum or integrate to one.

Probability Distributions

- Let A ⊂ R and let Prob(x ∈ A) denote the probability that x will belong to A.
- *Def.* The *distribution function* of a random variable x is a function defined by

 $F(x') \equiv Prob(x \leq x'), x' \in R.$

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Properties

P.1 F is nondecreasing in x.

P.2 $\lim_{x\to\infty} F(x) = 1$ and $\lim_{x\to-\infty} F(x) = 0$.

- P.3 F is continuous from the right.
- P.4 For all x', Prob(x > x') = 1 F(x').

Properties

P.5 For all x' and x" such that x" > x', $Prob(x' < x \le x'') = F(x'') - F(x')$.

P.6 For all x', $Prob(x < x') = \lim_{x \to x^{-}} F(x)$.

P.7 For all x', Prob(x=x') = $\lim_{x \to x^{+}} F(x) - \lim_{x \to x^{-}} F(x)$.

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Discrete Random Variables

 If the random variable can assume only a finite number or a countable infinite set of values, it is said to be a discrete random variable.

Key Properties

P.1 $Prob(x = x') \equiv f(x') \ge 0$. (f is called the *probability mass function* or the *probability function*.)

P.2
$$\sum_{i=1}^{\infty} f(x_i) = \sum_{i=1}^{\infty} \Pr{ob(x = x_i)} = 1.$$

P.3 Prob $(x \in A) = \sum_{x_i \in A} f(x_i)$.

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Examples

Example: #1 Consider the random variable associated with 2 tosses of a fair coin. The possible values for the #heads x are $\{0, 1, 2\}$. We have that f(0) = 1/4, f(1) = 1/2, and f(2) = 1/4.



Examples

#2 A single toss of a fair die.

f(x) = 1/6 if $x_i = 1,2,3,4,5,6$.

 $F(x_i) = x_i/6.$

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Continuous Random Variables and their Distributions

Def. A random variable x has a continuous distribution if there exists a nonnegative function f

defined on R such that for any interval A of R

Prob $(x \in A) = \int_{x \in A} f(x) dx$.

The function f is called the *probability density function* of x and the domain of f is called the *support* of the random variable x.

Properties of f

P.1 $f(x) \ge 0$, for all x.

P.2
$$\int_{-\infty}^{+\infty} f(x) dx = 1.$$

P.3 If dF/dx exists, then dF/dx = f(x), for all x.

In terms of geometry F(x) is the area under f(x) for $x' \le x$.

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Example

Example: The uniform distribution on [a,b].

$$f(x) = \begin{cases} 1/(b-a), \text{ if } x \in [a,b] \\ 0, \text{ otherwise} \end{cases}$$

Note that F is given by

$$F(x) = \int_{a}^{x} [1/(b-a)] dx = \frac{1}{(b-a)} x |_{a}^{x} = \frac{-a}{(b-a)} + \frac{1}{(b-a)} x.$$

Also,

$$\int_{a}^{b} f(x) dx = \int_{a}^{b} [1 / (b-a)] dx = \frac{1}{(b-a)} x |_{a}^{b} = \frac{-a}{(b-a)} + \frac{b}{(b-a)} = 1.$$

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Discrete Joint Distributions

• Let the two random variables x and y have a joint probability function

 $f(x_{i}', y_{i}') = Prob(x_{i} = x_{i}' and y_{i} = y_{i}').$

Properties of Prob Function

P.1 $f(x_i, y_i) \ge 0$.

P.2 Prob($(x_i, y_i) \in A$) = $\sum_{(x_i, y_i) \in A} f(x_i, y_i)$.

P.3 $\sum_{(x_i,y_i)} f(x_i,y_i) = 1.$

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The Distribution Function Defined

$$\begin{split} F(x_{i}^{'}, y_{i}^{'}) &= Prob(x_{i} \leq x_{i}^{'} \text{ and } y_{i} \leq y_{i}^{'}) = \sum_{(x_{i}^{'}, y_{i}^{'}) \in L} f(x_{i}^{'}, y_{i}^{'}), \text{ where} \\ L &= \{(x_{i}^{'}, y_{i}^{'}) : x_{i} \leq x_{i}^{'} \text{ and } y_{i} \leq y_{i}^{'}\}. \end{split}$$

Marginal Prob and Distribution Functions

- The marginal probability function associated with x is given by $f_1(x_j) \equiv Prob(x = x_j) = \sum_{y_i} f(x_{-j}, y_{-i})$
- The marginal probability function associated with y is given by f₂(y_j) ≡ Prob(y = y_j) = ∑_{x_i} f (x_i, y_j)

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Marginal distribution functions

• The marginal distribution function of x is given by

 $F_1(x_j) = \operatorname{Prob}(x_i \le x_j) = \lim_{y_i \to \infty} \operatorname{Prob}(x_i \le x_j \text{ and } y_i \le y_j) = \lim_{y_i \to \infty} F(x_j, y_j).$

• Likewise for y, the marginal distribution function is

$$F_2(y_j) = \lim_{x_j \to \infty} F(x_j, y_j).$$

Example

An example. Let x and y represent random variables representing whether or not two different stocks will increase or decrease in price. Each of x and y can take on the values 0 or 1, where a 1 means that its price has increased and a 0 means that it has decreased. The probability function is described by

 $f(1,1) = .50 \quad f(0,1) = .35 \quad f(1,0) = .10 \quad f(0,0) = .05.$

Answer each of the following questions.

- a. Find F(1,0) and F(0,1). F(1,0) = .1 + .05 = .15. F(0,1) = .35 + .05 = .40.
- b. Find $F_1(0) = \lim_{y \to 1} F(0,y) = F(0,1) = .4$
- c. Find $F_2(1) = \lim_{x \to 1} F(x,1) = F(1,1) = 1$.
- d. Find $f_1(0) = \sum_{y} f(0, y) = f(0, 1) + f(0, 0) = .4$.

e. Find
$$f_1(1) = \sum_{y} f(1, y) = f(1, 1) + f(1, 0) = .6$$

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Conditional Distributions

 After a value of y has been observed, the probability that a value of x will be observed is given by

 $Prob(x = x_i | y = y_i) = \frac{Prob(x = x_i \& y = y_i)}{Prob(y = y_i)} .$

• The function

$$g_1(\mathbf{x}_i \mid \mathbf{y}_i) \equiv \frac{\mathbf{f}(\mathbf{x}_i, \mathbf{y}_i)}{\mathbf{f}_2(\mathbf{y}_i)}$$

is called the *conditional probability function of x*, given y. $g_2(y_i | x_i)$ is defined analogously.

Properties of Conditional Probability Functions

 $(i) \hspace{0.2cm} g_{1}(x_{i} \mid y_{i}) \geq 0.$

(ii) $\sum_{x_i} g_1(x_i \mid y_i) = \sum_{x_i} f(x_i,y_i) / \sum_{x_i} f(x_i,y_i) = 1.$

((i) and (ii) hold for $g_2(y_i \mid x_i)$)

(iii) $f(x_i, y_i) = g_1(x_i \mid y_i)f_2(y_i) = g_2(y_i \mid x_i)f_1(x_i).$

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Conditional Distribution Functions

$$G_1(x_i | y_i) = \sum_{x \le x_i} f(x_i, y_i) / f_2(y_i),$$

$$G_2(\mathbf{y}_i \mid \mathbf{x}_i) = \sum_{\mathbf{y} \leq \mathbf{y}_i} \mathbf{f}(\mathbf{x}_i, \mathbf{y}_i) / \mathbf{f}_1(\mathbf{x}_i).$$

The stock price example revisited

a. Compute $g_1(1 \mid 0) = f(1,0)/f_2(0)$. We have that $f_2(0) = f(0,0) + f(1,0) = .05 + .1 = .15$. Further f(1,0) = .1. Thus, $g_1(1 \mid 0) = .1/.15 = .66$. b. Find $g_2(0 \mid 0) = f(0,0)/f_1(0) = .05/.4 = .125$. Here $f_1(0) = \sum_{y_i} f(0, y_i) = f(0,0) + f(0,1) = .05 + .35$

= .4.

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Continuous Joint Distributions

The random variables x and y have a continuous joint distribution if there exists a nonnegative function f defined on R² such that for any A ⊂ R²

Prob((x,y)
$$\in A$$
) = $\iint_A f(x,y) dx dy$.

• f is called the joint probability density function of x and y.

Properties of f

• f satisfies the usual properties:

P.1 $f(x,y) \ge 0$. P.2 $\int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x,y) dx dy = 1$.

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Distribution function

$$F(x',y') = \operatorname{Prob}(x \le x' \text{ and } y \le y') = \int_{-\infty}^{y'} \int_{-\infty}^{x'} f(x,y) dx dy.$$

If F is twice differentiable, then we have that $f(x,y) = \partial^2 F(x,y) / \partial x \partial y.$

Marginal Density and Distribution Functions

• The marginal density and distribution functions are defined as follows:

a. $F_1(x) = \lim_{y \to \infty} F(x,y)$ and $F_2(y) = \lim_{x \to \infty} F(x,y)$. (marginal distribution functions)

b. $f_1(x) = \int_y f(x,y) dy$ and $f_2(y) = \int_x f(x,y) dx$.

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Example

Let f(x,y) = 4xy for $x, y \in [0,1]$ and 0 otherwise.

a. Check to see that $\int_{0}^{1} \int_{0}^{1} 4xydxdy = 1$.

b. Find F(x',y'). Clearly, F(x',y') = 4 $\int_{0}^{y'} \int_{0}^{x'} xy dx dy = (x')^2 (y')^2$. Note also that $\partial^2 F/\partial x \partial y = 4xy = 4xy = 4xy$

f(x,y).

c. Find $F_1(x)$ and $F_2(y)$. We have that

 $F_1(x) = \lim_{y \to 1} x^2 y^2 = x^2.$

Using similar reasoning, $F_2(y) = y^2$.

d. Find $f_1(x)$ and $f_2(y)$.

$$f_1(x) = \int_{0}^{1} f(x,y)dy = 2x$$
 and $f_2(x) = \int_{0}^{1} f(x,y)dx = 2y$.

Conditional Density

• We have

The conditional density function of x, given that y is fixed at a particular value is given by

 $g_1(x | y) = f(x,y)/f_2(y).$

Likewise, for y we have

 $g_2(y | x) = f(x,y)/f_1(x).$

It is clear that $\int g_1(x | y) dx = 1$.

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Conditional Distribution Functions

• We have

The conditional distribution functions are given by

$$\begin{split} G_1(x' \mid y) &= \int\limits_{-\infty}^{x'} g_1(x \mid y) dx, \\ G_2(y' \mid x) &= \int\limits_{-\infty}^{y'} g_2(y \mid x) dy. \end{split}$$

Example

Let us revisit example #2 above. We have that f = 4xy with $x, y \in (0,1)$.

 $g_1(x \mid y) = 4xy/2y = 2x$ and $g_2(y \mid x) = 4xy/2x = 2y$.

Moreover,

$$G_1(x' | y) = 2 \int_{0}^{x'} x \, dx = 2 \frac{(x')^2}{2} = (x')^2.$$

By symmetry. $G_2(y | x) = (y')^2$. It turns out that in this example, x and y are independent random variables, because the conditional distributions do not depend on the other random variable.

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Independent Random Variables

Def. The random variables $(x_1, ..., x_n)$ are said to be independent if for any n sets of real numbers

A_i, we have

 $Prob(x_1 \in A_1 \And x_2 \in A_2 \And ... \And x_n \in A_n) = Prob(x_1 \in A_1) Prob(x_2 \in A_2) \bullet \bullet \bullet Prob(x_n \in A_n).$

Results on Independence

 The random variables x and y are independent iff

 $F(x,y) = F_1(x)F_2(y)$ or

 $f(x,y) = f_1(x)f_2(y).$

• Further, iff x and y are independent, then

 $g_1(x \mid y) = f(x,y)/f_2(y) = f_1(x)f_2(y)/f_2(y) = f_1(x).$

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Extensions

- The notion of a joint distribution can be extended to any number of random variables.
- The marginal and conditional distributions are easily extended to this case.
- Let $f(x_1,...,x_n)$ represent the joint density.

Extensions

The marginal density for the ith variable is given by

 $f_i(x_i) = \int \dots \int f(x_1, \dots, x_n) dx_1 \dots dx_{i-1} dx_{i+1} \dots dx_n.$

The conditional density for say x₁ given x₂,...,x_n is

 $g_1(x_1 | x_2,...,x_n) = f(x_1,...,x_n) / \int f(x_1,...,x_n) dx_1.$

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Summary Measures of Probability Distributions

- Summary measures are scalars that convey some aspect of the distribution. Because each is a scalar, all of the information about the distribution cannot be captured. In some cases it is of interest to know multiple summary measures of the same distribution.
- There are two general types of measures.

a. Measures of central tendency: Expectation, median and mode

b. measures of dispersion: Variance

Expectation

• The *expectation of a random variable x* is given by

 $E(x) = \sum x_i f(x_i)$ (discrete)

 $E(x) = \int xf(x)dx$. (continuous)

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Examples

#1. A lottery. A church holds a lottery by selling 1000 tickets at a dollar each. One winner wins\$750. You buy one ticket. What is your expected return?

E(x) = .001(749) + .999(-1) = .749 - .999 = -.25.

The interpretation is that if you were to repeat this game infinitely your long run return would be - .25.

#2. You purchase 100 shares of a stock and sell them one year later. The net gain is x_i . The distribution is given by. (-500, .03), (-250, .07), (0,.1), (250, .25),(500, .35), (750, .15), and (1000, .05).

E(x) = \$367.50

Examples

#3. Let f(x) = 2x for $x \in (0,1)$ and = 0, otherwise. Find E(x).

$$E(x) = \int_{0}^{1} 2x^{2} dx = 2/3.$$

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Properties of E(x)

P.1 Let g(x) be a function of x. Then E(g(x)) is given by

 $E(g(x)) = \Sigma g(x_i) f(x_i)$ (discrete)

 $E(g(x)) = \int g(x)f(x) dx$. (continuous)

P.2 If k is a constant, then E(k) = k.

P.3 Let a and b be two arbitrary constants. Then E(ax+b) = aE(x)+b.

Properties of E(x)

P.4 Let $x_1, ..., x_n$ be n random variables. Then $E(\Sigma x_i) = \Sigma E(x_i)$.

P.5 If there exists a constant k such that $Prob(x \ge k) = 1$, then $E(x) \ge k$. If there exists a constant

k such that $Prob(x \le k) = 1$, then $E(x) \le k$.

P.6 Let $x_1, ..., x_n$ be n independent random variables. Then $E(\prod_{i=1}^n x_i) = \prod_{i=1}^n E(x_i)$.

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Median

- Def. If Prob(x ≤ m) ≥ .5 and Prob(x ≥ m)
 ≥ .5, then m is called a *median*.
- a. The continuous case

$$\int_{-\infty}^{m} f(x) dx = \int_{m}^{+\infty} f(x) dx = .5.$$

b. In the discrete case, m need not be unique. Example: $(x_1,f(x_1))$ given by (6,.1), (8,.4), (10, .3), (15, .1), (25, .05),(50, .05). In this case, m = 8 or 10.

Mode

- Def. The mode is given by m_o = argmax f(x).
- A mode is a maximizer of the density function. It need not be unique.

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A Summary Measure of Dispersion: The Variance

- In many cases the mean the mode or the median are not informative.
- In particular, two distributions with the same mean can be very different distributions. One would like to know how common or typical is the mean. The variance measures this notion by taking the expectation of the squared deviation about the mean.

Variance

- Def. For a random variable x, the variance is given by E[(x-μ)²], where μ = E(x).
- The variance is also written as Var(x) or as σ². The square root of the variance is called the *standard deviation* of the distribution. It is written as σ.

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Illustration

	Jan	Feb	Mar	Apr	Мау	Jun	Jul	Aug	Sep	Oct	Nov	Dec	E(x)	Var(x)
Lawrence	28	34	46	57	66	75	80	78	70	58	46	34	56	334
Santa Barbara	52	54	55	57	58	62	65	67	66	62	57	52	58.91667	28.62879

Computation: Examples

a. For the discrete case, $Var(x) = \Sigma (x_i - \mu)^2 f(x_i)$. As an example, if $(x_i, f(x_i))$ are given by (0, .1), (500, .8), and (1000, .1). We have that E(x) = 500.

 $Var(x) = (0-500)^{2}(.1) + (500 - 500)^{2}(.8) + (1000 - 500)^{2}(.1) = 50,000.$

b. For the continuous case, $Var(x) = \int (x-\mu)^2 f(x) dx$. Consider the example above where f = 2x with $x \in (0,1)$. From above, E(x) = 2/3. Thus,

Var(x) =
$$\int_{0}^{1} (x - 2/3)^{2} 2x \, dx = 1/18.$$

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Properties of Variance

- P.1 Var(x) = 0 iff there exists a c such that Prob(x = c) = 1.
- P.2 For any constants a and b, $Var(ax +b) = a^2 Var(x)$.

P.3 Var(x) = $E(x^2) - [E(x)]^2$.

- P.4 If x_i , i = 1, ..., n, are independent, then $Var(\Sigma x_i) = \Sigma Var(x_i)$.
- P.5 If x_i are independent, i = 1, ..., n, then $Var(\Sigma a_i x_i) = \Sigma a_i^2 Var(x_i)$.

A remark on moments

- Var (x) is sometimes called the second moment about the mean, with E(x-μ) = 0 being the first moment about the mean.
- Using this terminology, E(x-μ)³ is the *third* moment about the mean. It can give us information about the skewedness of the distribution. E(x-μ)⁴ is the *fourth moment* about the mean and it can yield information about the modes of the distribution or the peaks (kurtosis).

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Moments of Conditional and Joint Distributions

Given a joint probability density function $f(x_1, ..., x_n)$, the expectation of a function of the n

variables say $g(x_1, ..., x_n)$ is defined as

 $E(g(x_1, \ldots, x_n)) = \int \bullet \bullet \bullet \int g(x_1, \ldots, x_n) f(x_1, \ldots, x_n) dx_1 \bullet \bullet \bullet dx_n.$

If the random variables are discrete, then we would let $x^i = (x_1^{i_1}, ..., x_n^{i_n})$ be the i^{th} observation and write

 $E(g(x_1, \ldots, x_n)) = \Sigma g(x^i) f(x^i).$

Unconditional expectation of a joint distribution

 Given a joint density f(x,y), E(x) is given by

$$E(x) = \int_{-\infty}^{+\infty} xf_1(x)dx = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xf(x,y)dxdy.$$

• Likewise, E(y) is

$$E(y) = \int_{-\infty}^{+\infty} yf_2(y)dy = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} yf(x,y)dxdy.$$

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Conditional Expectation

 The conditional expectation of x given that x and y are jointly distributed as f(x,y) is defined by (I will give definitions for the continuous case only. For the discrete case, replace integrals with summations)

$$\mathbf{E}(\mathbf{x} \mid \mathbf{y}) = \int_{-\infty}^{+\infty} \mathbf{x} \mathbf{g}_{\mathbf{i}}(\mathbf{x} \mid \mathbf{y}) \, \mathrm{d}\mathbf{x}$$

Conditional Expectation

• Further the conditional expectation of y given x is defined analogously as

$$E(y \mid x) = \int_{-\infty}^{+\infty} yg_2(y \mid x) \, dy$$

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Conditional Expectation

 Note that E(E(x | y)) = E(x). To see this, compute

 $E(E(x \mid y)) = \int_{-\infty}^{+\infty} \left[\int_{-\infty}^{+\infty} xg_1(x \mid y) dx \right] f_2 dy = \int_{-\infty}^{+\infty} \left\{ \int_{-\infty}^{+\infty} x[f(x,y)/(f_2)] dx \right\} f_2 dy$ $= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xf(x,y) dx dy,$

and the result holds.

Covariance.

• *Covariance* is a moment reflecting direction of movement of two variables. It is defined as

 $Cov(x,y) = E[(x-\mu_x)(y-\mu_y)].$

 When this is large and positive, then x and y tend to be both much above or both much below their respective means at the same time. Conversely when it is negative.

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Computation of Cov

. Computation of the covariance. First compute

 $(x-\mu_x)(y-\mu_y) = xy - \mu_x y - \mu_y x + \mu_x \mu_y.$

Taking E,

 $E(xy) - \mu_x \mu_y - \mu_x \mu_y + \mu_x \mu_y = E(xy) - \mu_x \mu_y.$

Thus, Cov(xy) = E(xy) - E(x)E(y). If x and y are independent, then E(xy) = E(x)E(y) and Cov(xy) = 0.